

## EMPTY NON-CONVEX AND CONVEX FOUR-GONS IN RANDOM POINT SETS

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### Abstract

Let  $\mathcal{S}$  be a set of  $n$  points distributed uniformly and independently in a convex, bounded set in the plane. A four-gon is called empty if it contains no points of  $\mathcal{S}$  in its interior. We show that the expected number of empty non-convex four-gons with vertices from  $\mathcal{S}$  is  $12n^2 \log n + o(n^2 \log n)$  and the expected number of empty convex four-gons with vertices from  $\mathcal{S}$  is  $\Theta(n^2)$ .

### 1. Introduction

Sylvester's famous four-point problem asks for the probability that four points chosen at random inside a convex set  $\mathcal{K}$  in the plane form the vertices of a non-convex (also called reentrant) quadrilateral. The answer to this question depends on the shape of  $\mathcal{K}$ , see the historical account of Pfeifer [16]. Almost a century ago, Blaschke [10] proved that this probability is maximal if  $\mathcal{K}$  is a triangle, and minimal if  $\mathcal{K}$  is an ellipse. Actually, this problem is equivalent to finding the expected triangle area  $\mathbb{E}_\Delta$  of three points chosen at random from  $\mathcal{K}$ .

More precisely, the probability that a four-gon is non-convex is four times  $\mathbb{E}_\Delta$  divided by the area of  $\mathcal{K}$ .

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The question studied here is related: A set  $\mathcal{S}$  of  $n$  points is distributed at random inside  $\mathcal{K}$ . What is the expected number of empty four-gons with vertices from  $\mathcal{S}$ ?

Figure 1 shows an empty non-convex and an empty convex four-gon in  $\mathcal{S}$ . The probability that a four-gon is empty depends on its area, and we will make use of the density function of the triangle area of three points chosen at random from  $\mathcal{K}$ , see Henze [13]. Also, the expected numbers of empty non-convex and empty convex four-gons are different. Further, we observe that the asymptotic order of magnitude of the number of empty four-gons does not depend on the shape of the convex set. Such a behaviour was also obtained in related works: Valtr [17] proved that the expected number of empty triangles in  $\mathcal{S}$  is at most  $2n^2 - 2n$  and at least  $2n^2 - o(n^2)$ , independent of the shape of  $\mathcal{K}$ .

Bárány et al. [8] proved that in expectation  $\mathcal{S}$  contains two points which both appear in  $\Omega(n/\log n)$  empty triangles.

Bárány and Füredi [7] proved that the number of empty simplices in random point sets of  $n$  points distributed uniformly and independently in a convex, bounded set in  $R^d$  is at most  $K \binom{n}{d}$ , for some constant  $K$ .

Balogh et al. [5] showed that the expected number of vertices of the largest empty convex polygon in  $\mathcal{S}$  is  $\Theta\left(\frac{\log n}{\log \log n}\right)$ .

Sylvester's four-point problem was generalized by asking for the probability that  $n$  points chosen at random inside  $\mathcal{K}$  are in convex position [6, 15, 18, 19]. In this case the answer depends on the shape of  $\mathcal{K}$ .

A lot of research has been done to determine the minimum number  $f_k(n)$  of empty convex  $k$ -gons among all sets of  $n$  points in general position in the plane (not only random point sets).

For the case of empty triangles, Katchalski and Meir [14] showed that  $f_3(n)$  is of order  $\Theta(n^2)$ . Later, this bound was refined [2, 7, 9, 11, 12, 17]; the currently best bounds are  $n^2 - \frac{32n}{7} + \frac{22}{7} \leq f_3(n) \leq 1.6196n^2 + o(n^2)$ .

Concerning empty convex four-gons, Bárány and Füredi [7] established that  $f_4(n)$  is of order  $\Theta(n^2)$ , and the currently best bounds on  $f_4(n)$  are  $\frac{n^2}{2} - \frac{9}{4}n - o(n) \leq f_4(n) \leq 1.9397n^2 + o(n^2)$ , see [2, 9].

Research mainly focussed on empty convex polygons. Only recently the number of convex and non-convex polygons in point sets has been studied [1, 3, 4].

In [1] it was shown that every set of  $n$  points in general position in the plane determines at least  $\frac{5n^2}{2} - \Theta(n)$  empty four-gons and a point set with only  $O(n^{5/2} \log n)$  empty four-gons is given. Our result improves the latter bound to  $O(n^2 \log n)$ .

Let us state the results more formally. All logarithms in this paper are natural logarithms. Throughout this paper let  $\mathcal{S}$  be a set of  $n$  points distributed uniformly and independently in a convex, bounded set  $\mathcal{K}$  of unit area in the plane. Since with probability 1 the  $n$  points will be in gen-

eral position (no three points are collinear), we may and will assume this throughout the paper. All asymptotics in this paper are w.r.t. the number of points  $n$ , that is, when  $n \rightarrow \infty$ . As our results are asymptotic, we may ignore also rounding issues throughout the paper, that is, if for some constant  $c > 0$ ,  $cn$  is not an integer, depending on the context, we may and will consider  $\lfloor cn \rfloor$  or  $\lceil cn \rceil$  without changing the results.

A four-gon whose vertices are from  $\mathcal{S}$  is empty if it contains no other point from  $\mathcal{S}$  in its interior.

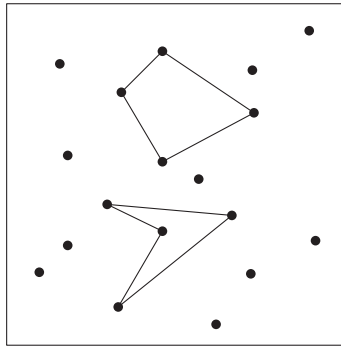


Fig. 1. An empty non-convex four-gon and an empty convex four-gon in  $\mathcal{S}$

Denote by  $N_4$  the random variable that counts the number of empty non-convex four-gons with vertices from  $\mathcal{S}$ .

Our main result is the following:

**THEOREM 1.1.**  $\mathbb{E}[N_4] = 12n^2 \log n + o(n^2 \log n)$ .

Denote by  $C_4$  the random variable that counts the number of empty convex four-gons with vertices from  $\mathcal{S}$ .

Complementing Theorem 1.1, we prove the following result (the lower bound was already obtained in [7]. There, also another construction with  $O(n^2)$  empty convex four-gons is given.)

**THEOREM 1.2.**  $\mathbb{E}[C_4] = \Theta(n^2)$ .

## 2. Proof of Theorem 1.1

The proof of Theorem 1.1 is implied by the following lemmas, for which we need some definitions. Fix three points  $p_a, p_b, p_c$  and denote by  $\Delta(p_a, p_b, p_c)$  the triangle spanned by them. Let  $\mathcal{P}$  be a set of  $k \geq 1$  points distributed uniformly and independently in  $\Delta(p_a, p_b, p_c)$ . Denote by  $\mathcal{E}_{p_a p_b, p_b p_c}^{\mathcal{P}}$  the event that  $\mathcal{P} \cup \{p_a, p_b, p_c\}$  contains an empty non-convex four-gon with  $p_a p_b$  and  $p_b p_c$  among its edges.

LEMMA 2.1.  $\mathbb{P}(\mathcal{E}_{p_a p_b, p_b p_c}^{\mathcal{P}}) = \frac{2}{k+1}$ .

PROOF. First observe that the points  $\{p_a, p_b, p_c\}$  together with a fourth point  $p_d \in \mathcal{P}$  form an empty non-convex four-gon with  $p_a p_b$  and  $p_b p_c$  among its edges if and only if the triangle  $\Delta(p_a, p_d, p_c)$  contains  $\mathcal{P} \setminus \{p_d\}$  in its interior. We now determine the distribution of the height  $h_d$ , which is the distance from  $p_d$  to the segment  $p_a p_c$ . Denote by  $h$  the height of  $\Delta(p_a, p_b, p_c)$  with respect to the edge  $p_a p_c$ . Let  $\ell$  be the segment parallel to the edge  $p_a p_c$ , at distance  $h_d$  from this edge, and with endpoints on the edges  $p_a p_b$  and  $p_b p_c$  respectively. Assume w.l.o.g. that  $\ell$  is a horizontal line with  $p_a$  and  $p_c$  below it. Define then  $\Delta(\ell, p_b)$  to be the triangle coming from the intersection of  $\Delta(p_a, p_b, p_c)$  and all points lying on or above  $\ell$ , as shown in Figure 2.

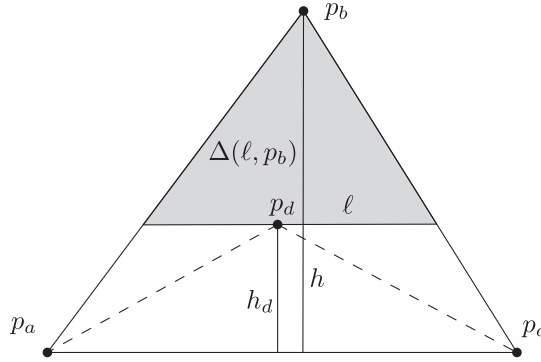


Fig. 2. The triangles  $\Delta(p_a, p_b, p_c)$ ,  $\Delta(p_a, p_d, p_c)$  and  $\Delta(\ell, p_b)$

Since only relative heights between  $h$  and  $h_d$  matter, we may assume w.l.o.g. that  $h = 1$ .

By the intercept theorem we have  $\frac{|p_a p_c|}{h} = \frac{|\ell|}{h - h_d}$ , where  $|p_a p_c|$  and  $|\ell|$  are the lengths of the segments  $p_a p_c$  and  $\ell$ .

It follows that

$$\frac{\text{area}(\Delta(\ell, p_b))}{\text{area}(\Delta(p_a, p_b, p_c))} = \frac{\frac{|\ell|(h - h_d)}{2}}{\frac{|p_a p_c|h}{2}} = (1 - h_d)^2.$$

Hence, the distribution function  $F_{h_d}$  for the height  $h_d$  satisfies  $F_{h_d}(x) = \mathbb{P}(h_d \leq x) = 1 - (1 - x)^2$  and the density of the height  $h_d$  is  $f_{h_d}(x) = 2 - 2x$  for  $x \in [0, 1]$ .

Fix any  $p \in \mathcal{P} \setminus \{p_d\}$ . Since the points are distributed uniformly at random inside  $\Delta(p_a, p_b, p_c)$ , we have

$$\mathbb{P}(p \in \Delta(p_a, p_d, p_c)) = \frac{\text{area}(\Delta(p_a, p_d, p_c))}{\text{area}(\Delta(p_a, p_b, p_c))} = h_d.$$

Therefore, integrating over all possible heights  $0 \leq h_d \leq 1$ ,

$$\mathbb{P}(\mathcal{P} \setminus \{p_d\} \in \Delta(p_a, p_d, p_c)) = \int_0^1 x^{k-1}(2-2x) dx = \frac{2}{k(k+1)}.$$

As there are  $k$  choices for the point  $p_d$ , by a union bound, we have

$$\mathbb{P}(\mathcal{E}_{p_a p_b, p_b p_c}^{\mathcal{P}}) \leq \frac{2}{k+1}.$$

On the other hand, there is at most one point  $p_d \in \mathcal{P}$  such that  $\mathcal{P} \setminus \{p_d\} \in \Delta(p_a, p_d, p_c)$ : indeed, if this were true for another point  $p_e \neq p_d$ , then  $p_d \notin \Delta(p_a, p_e, p_c)$ , contradicting the assumption. Hence,

$$\mathbb{P}((\mathcal{P} \setminus \{p_d\} \in \Delta(p_a, p_d, p_c)) \wedge (\mathcal{P} \setminus \{p_e\} \in \Delta(p_a, p_e, p_c))) = 0,$$

and thus

$$\begin{aligned} \mathbb{P}(\mathcal{E}_{p_a p_b, p_b p_c}^{\mathcal{P}}) &= \bigcup_{p_d \in \mathcal{P}} \mathbb{P}(\mathcal{P} \setminus \{p_d\} \in \Delta(p_a, p_d, p_c)) \\ &= \sum_{p_d \in \mathcal{P}} \mathbb{P}(\mathcal{P} \setminus \{p_d\} \in \Delta(p_a, p_d, p_c)) = \frac{2}{k+1}. \quad \square \end{aligned}$$

For the next lemma, we need one more definition. Let  $T_k$  denote the random variable that counts the number of triangles with vertices in  $\mathcal{S}$  containing exactly  $k \geq 0$  points from  $\mathcal{S}$  in its interior.

LEMMA 2.2. *For any  $k = k(n) \geq 0$ ,  $\mathbb{E}[T_k] \leq 2n^2 - 2n$ .*

PROOF. Let  $f_{\text{area}(\Delta)}(v)$  denote the density function for the area  $v$  of a triangle  $\Delta$  formed by three points chosen uniformly and independently in the bounded, convex set  $\mathcal{K}$ . This density function has been studied in [13]; also see references therein. From the fact that  $f_{\text{area}(\Delta)}(0) = 12$  and that  $f_{\text{area}(\Delta)}(v)$  is a monotonically decreasing function, we have that for any area  $v \geq 0$ ,  $f_{\text{area}(\Delta)}(v) \leq 12$ .

We remark that it also follows from Lemma 5.1 of [7] that  $f_{\text{area}(\Delta)}(v)$  is bounded from above by a constant.

Fix now three points  $p_a, p_b, p_c \in \mathcal{S}$ , let as before  $\Delta(p_a, p_b, p_c)$  be the triangle spanned by them, and let  $\text{int}(\Delta(p_a, p_b, p_c))$  denote the interior of this triangle. Denote also for  $x, y > 0$  by  $\beta(x, y) = \int_0^1 t^{x-1}(1-t)^{y-1} dt$  the beta function of  $x$  and  $y$ .

Integrating over all possible areas  $v$  of the triangle  $\Delta(p_a, p_b, p_c)$ , we obtain

$$\begin{aligned} \mathbb{P}(|\text{int}(\Delta(p_a, p_b, p_c)) \cap \mathcal{S}| = k) &= \int_0^1 \binom{n-3}{k} v^k (1-v)^{n-3-k} f_{\text{area}(\Delta)}(v) dv \\ &\leq 12 \int_0^1 \binom{n-3}{k} v^k (1-v)^{n-3-k} dv \\ &= 12 \binom{n-3}{k} \beta(k+1, n-k-2) \\ &= 12 \binom{n-3}{k} \frac{k!(n-k-3)!}{(n-2)!} = \frac{12}{n-2}. \end{aligned}$$

Hence, by linearity of expectation, for any  $k = k(n) \geq 0$ ,

$$\mathbb{E}[T_k] \leq \binom{n}{3} \frac{12}{n-2} = 2n^2 - 2n. \quad \square$$

REMARK. The special case  $k = 0$  of Lemma 2.2 was also proved by Valtr in [17].

LEMMA 2.3. *For every  $\varepsilon > 0$  there exists some  $\alpha = \alpha(\varepsilon) > 0$  such that  $\mathbb{E}[T_k] \geq (2 - \varepsilon)n^2$  for any  $k = 0, 1, \dots, \alpha n$ .*

PROOF. The density function  $f_{\text{area}(\Delta)}(v)$  for the area of the triangle  $\Delta = \Delta(p_a, p_b, p_c)$ , formed by three points  $p_a, p_b, p_c$  from  $\mathcal{S}$ , satisfies  $f_{\text{area}(\Delta)}(0) = 12$  and is then strictly monotonically decreasing. In particular, for every small  $\varepsilon > 0$  there exists  $v_0 > 0$  such that  $f_{\text{area}(\Delta)}(v_0) = 12 - \varepsilon$ . We define  $\alpha = 0.6v_0$ .

$$\begin{aligned} \mathbb{P}(|\text{int}(\Delta) \cap \mathcal{S}| = k) &= \int_0^1 \binom{n-3}{k} v^k (1-v)^{n-3-k} f_{\text{area}(\Delta)}(v) dv \\ &\geq (12 - \varepsilon) \int_0^{v_0} \binom{n-3}{k} v^k (1-v)^{n-3-k} dv. \end{aligned}$$

As in the proof of Lemma 2.2 we have

$$\int_0^1 \binom{n-3}{k} v^k (1-v)^{n-3-k} dv = \frac{1}{n-2}.$$

We will show that  $\mathbb{P}(|\text{int}(\Delta) \cap \mathcal{S}| = k) \geq \frac{12-\varepsilon}{n-2} - o\left(\frac{1}{n}\right)$ , for  $k \leq \alpha n$ . To this end, it is sufficient to show that

$$(1) \quad \int_{v_0}^1 \binom{n-3}{k} v^k (1-v)^{n-3-k} dv = o\left(\frac{1}{n}\right).$$

Computing the derivative of the function  $g(v) := v^k (1-v)^{n-3-k}$  in  $[v_0, 1]$ , we see that  $g'(v) \leq 0$ , implying that in  $[v_0, 1]$ ,  $g(v)$  is maximized at  $v = v_0$ .

Thus,

$$\int_{v_0}^1 \binom{n-3}{k} v^k (1-v)^{n-3-k} dv \leq \binom{n-3}{k} v_0^k (1-v_0)^{n-3-k} (1-v_0).$$

It is easily verified that

$$\binom{n-3}{k} v_0^k (1-v_0)^{n-3-k} < \binom{n-3}{k+1} v_0^{k+1} (1-v_0)^{n-3-k-1}$$

holds for  $k < v_0(n-2) - 1$ .

If we can show that

$$\binom{n-3}{k} v_0^k (1-v_0)^{n-3-k} (1-v_0) = o\left(\frac{1}{n}\right)$$

holds for  $k = \alpha n$ , then it holds for all smaller values of  $k$  as well, and then also (1) holds for all  $k = 0, 1, \dots, \alpha n$ . Using  $\binom{n}{k} \leq \left(\frac{ne}{k}\right)^k$ , we get for some constant  $C > 0$

$$\begin{aligned} & \binom{n-3}{\alpha n} v_0^{\alpha n} (1-v_0)^{n-3-\alpha n} (1-v_0) \\ & \leq C \left(\frac{e}{0.6v_0}\right)^{0.6v_0 n} v_0^{0.6v_0 n} (1-v_0)^{(1-0.6v_0)n} \end{aligned}$$

$$\begin{aligned}
 &= C((e/0.6)^{0.6v_0}(1-v_0)^{1-0.6v_0})^n \\
 &= o\left(\frac{1}{n}\right),
 \end{aligned}$$

where the last line follows from the facts that

$$f(v_0) := \left(\frac{e}{0.6}\right)^{0.6v_0}(1-v_0)^{1-0.6v_0}$$

is monotone decreasing for  $v_0 \in [0, 1]$ , that  $v_0 > 0$ , and that  $f(0) = 1$ . Thus,

$$\begin{aligned}
 \mathbb{P}(|\text{int}(\Delta) \cap \mathcal{S}| = k) &\geq (12 - \varepsilon) \int_0^{v_0} \binom{n-3}{k} v^k (1-v)^{n-3-k} dv \\
 &\geq \frac{12 - \varepsilon}{n - 2} - o\left(\frac{1}{n}\right).
 \end{aligned}$$

As before, by linearity of expectation,

$$\mathbb{E}[T_k] \geq \binom{n}{3} \left( \frac{12 - \varepsilon}{n - 2} - o\left(\frac{1}{n}\right) \right) \geq (2 - \varepsilon)n^2$$

for any  $k = 0, 1, \dots, \alpha n$ , thus concluding the proof. □

We are now ready to prove Theorem 1.1.

**PROOF OF THEOREM 1.1.** Note that each triangle  $\Delta(p_a, p_b, p_c)$  with vertices  $p_a, p_b, p_c \in \mathcal{S}$  determines at most three empty non-convex four-gons such that  $p_a, p_b, p_c$  are the vertices on the boundary of the convex hull of the quadrilateral: indeed, any pair of edges from  $\{p_a p_b, p_b p_c, p_a p_c\}$  can be chosen and might possibly give rise to an empty non-convex four-gon. Let  $\mathcal{P} \subseteq \mathcal{S}$  denote the set of points in the interior of a triangle  $\Delta(p_a, p_b, p_c)$ . We classify the empty non-convex four-gons in  $\mathcal{S}$  according to the cardinality of  $\mathcal{P}$ . That is, we define  $N_{4,k}$ , for  $1 \leq k \leq n - 3$ , as the random variable that counts the number of empty non-convex four-gons in all triangles  $\Delta(p_a, p_b, p_c)$  with  $|\mathcal{P}| = k$  points of  $\mathcal{S}$  in its interior, such that the four-gons associated to a triangle  $\Delta(p_a, p_b, p_c)$  have  $p_a, p_b, p_c$  among their vertices. Then  $N_4 = \sum_{k=1}^{n-3} N_{4,k}$ .

Let  $X_{k,p_a p_b, p_b p_c}$  be the indicator random variable for the event that a triangle  $\Delta(p_a, p_b, p_c)$  with  $|\mathcal{P}| = k$  points of  $\mathcal{S}$  in its interior, contains an empty non-convex four-gon with  $p_a p_b$  and  $p_b p_c$  among its edges. Analogously,  $X_{k,p_b p_c, p_c p_a}$  and  $X_{k,p_c p_a, p_a p_b}$  are the indicator random variables for the events that  $\Delta(p_a, p_b, p_c)$  contains an empty non-convex four-gon



with  $p_b p_c$  and  $p_c p_a$ , respectively  $p_c p_a$  and  $p_a p_b$ , among its edges. For  $X_k = X_{k,p_a p_b, p_b p_c} + X_{k,p_b p_c, p_c p_a} + X_{k,p_c p_a, p_a p_b}$  we get from Lemma 2.1 that  $\mathbb{E}[X_k] = \frac{6}{k+1}$ .

We have

$$\mathbb{E}[N_{4,k} \mid T_k = t] = t\mathbb{E}[X_k].$$

Then

$$\begin{aligned} \mathbb{E}[N_{4,k}] &= \sum_{t=0}^{\binom{n}{3}} \mathbb{E}[N_{4,k} \mid T_k = t] \mathbb{P}(T_k = t) \\ &= \sum_{t=0}^{\binom{n}{3}} t \mathbb{E}[X_k] \mathbb{P}(T_k = t) = \mathbb{E}[X_k] \sum_{t=0}^{\binom{n}{3}} t \mathbb{P}(T_k = t) = \mathbb{E}[X_k] \mathbb{E}[T_k] \end{aligned}$$

and

$$\mathbb{E}[N_4] = \sum_{k=1}^{n-3} \mathbb{E}[N_{4,k}] = \sum_{k=1}^{n-3} \mathbb{E}[X_k] \mathbb{E}[T_k].$$

By Lemma 2.2,

$$\mathbb{E}[N_4] \leq (2n^2 - 2n) \sum_{k=1}^{n-3} \frac{6}{k+1} \leq (12n^2 - 12n) \log n.$$

By Lemma 2.3,

$$\mathbb{E}[N_4] \geq (2 - \varepsilon)n^2 \sum_{k=1}^{\alpha n} \frac{6}{k+1} \geq 6(2 - \varepsilon)n^2((\log \alpha n) - 1).$$

Since the latter inequality holds for every  $\varepsilon > 0$  and  $\alpha(\varepsilon)$  we conclude that

$$\mathbb{E}[N_4] \geq 12n^2 \log n - o(n^2 \log n).$$

This completes the proof of Theorem 1.1.  $\square$

**3. Proof of Theorem 1.2**

The following proof roughly follows the proof of Valtr [17] for the upper bound of  $2n^2 - 2n$  empty triangles in  $\mathcal{S}$ .

PROOF. For each convex four-gon we focus on its largest (interior) diagonal.

Let  $p_a, p_b \in \mathcal{S}$  and consider four-gons whose largest diagonal is  $p_a p_b$  of length  $|p_a p_b| = \ell$ . Note that, since  $\mathcal{S}$  is a set of points distributed in a bounded set,  $\ell$  is bounded from above by a constant  $D$ , which is the diameter of  $\mathcal{K}$ . Suppose w.l.o.g. that for the coordinates of  $p_a$  and  $p_b$  we have  $p_a = (\ell, \ell), p_b = (2\ell, \ell)$ . Consider the two axis-parallel rectangles  $R_1$  and  $R_2$  of width  $3\ell$  and height  $\ell$  whose left lower cornerpoints are  $(0, \ell)$ , and  $(0, 0)$ , respectively, as shown in Figure 3.

Observe that if  $p_a p_b$  is the largest diagonal of a convex four-gon with vertex set  $\{p_a, p_b, p_c, p_d\}$ , then it is necessary that the other two points  $p_c$  and  $p_d$  are in  $R_1$  and  $R_2$  (one in  $R_1$ , and one in  $R_2$ ): indeed, since the four-gon is convex, both diagonals are inside, and hence if one of the points  $p_c$  or  $p_d$  were outside the rectangles, as the segment  $p_c p_d$  has to cross the diagonal  $p_a p_b$ , its length were bigger than  $\ell$ , contradicting the fact that  $p_a p_b$  is the longest diagonal.

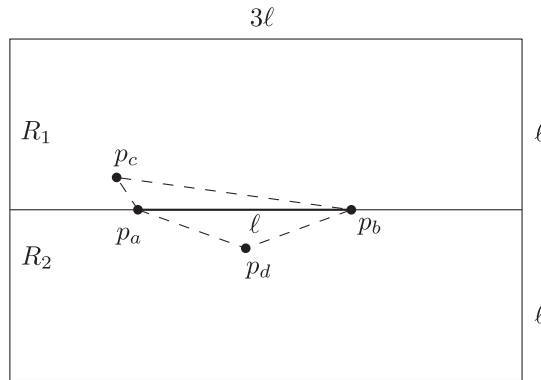


Fig. 3. The bounding box containing points  $p_c$  and  $p_d$  of a four-gon with largest diagonal  $p_a p_b$  and  $|p_a p_b| = \ell$

Now, fix 4 points  $p_a, p_b, p_c, p_d \in \mathcal{S}$  (ordered, only the order between  $p_a$  and  $p_b$  does not matter) and define by  $\mathcal{E}_{p_a p_b, p_c, p_d}$  the event that the 4 points  $p_a, p_b, p_c, p_d$  form an empty convex four-gon whose largest diagonal is formed by  $p_a p_b$ . Denote by  $f_{\text{len}}(x)$  the probability density function that the length of the edge between two randomly chosen points in  $\mathcal{K}$  is exactly  $x$ . Clearly,

$$f_{\text{len}}(x) \leq 2x\pi,$$

since the position of the first point is arbitrary, and the second point then has to be on the circumference of a ball of radius  $x$  centered at the first point. Denote also by  $F_{\text{hei}}(h)$  the probability that the  $y$ -coordinate of a randomly chosen point in  $\mathcal{K}$  is in  $[\ell, \ell + h]$ , and simultaneously the  $x$ -coordinate is in  $[0, 3\ell]$ . Let  $f_{\text{hei}}(h)$  be the corresponding probability density function of additionally having  $y$ -coordinate exactly  $\ell + h$ . We have

$$F_{\text{hei}}(h) \leq 3\ell h, \quad \text{and} \quad f_{\text{hei}}(h) \leq 3\ell.$$

Note that for  $h \leq \ell$ ,  $f_{\text{hei}}(h)$  corresponds to the probability density that a randomly chosen point from  $\mathcal{S} \setminus \{p_a, p_b\}$  is inside  $R_1$ , and at vertical distance  $h$  measured from  $p_a p_b$  (in fact, it is an upper bound, since depending on the value of  $\ell$  the rectangle might not be fully contained in  $\mathcal{K}$ ). By symmetry, this is also an upper bound for the probability density function to be at vertical distance  $h$  from  $p_a p_b$  and inside  $R_2$ . Call this vertical distance in both cases the *height* of a point.

Since we consider  $p_c$  and  $p_d$  to be an ordered pair, we consider for  $p_c$  only the probability density of all heights  $h_c$  in  $R_1$ , and for  $p_d$  only the probability density of all heights  $h_d$  in  $R_2$ . Assuming that  $p_c$  and  $p_d$  are at heights  $h_c$  and  $h_d$  in the corresponding rectangles, let  $\Delta(\ell, p_c)$  be the triangle with base edge  $p_a p_b$  and height  $h_c$ , going through the point  $p_c$ , and analogously for  $\Delta(\ell, p_d)$ . Note that for  $1 \leq \ell \leq D$ , we have  $h_c \leq \frac{2}{\ell}$  and  $h_d \leq \frac{2}{\ell}$ , as the areas of  $\Delta(\ell, p_c)$  and  $\Delta(\ell, p_d)$  are bounded by 1.

Also, by convexity, if all of  $p_a, p_b, p_c, p_d$  are inside  $\mathcal{K}$ , then also both triangles fall entirely into  $\mathcal{K}$ . We estimate the probability that there are no other points in  $\mathcal{S} \setminus \{p_a, p_b, p_c, p_d\}$  that fall into the triangles (with disjoint interiors)  $\Delta(\ell, p_c)$  and  $\Delta(\ell, p_d)$  of total area  $\frac{\ell(h_c + h_d)}{2}$ . Distinguishing the two cases  $0 \leq \ell < 1$  and  $1 \leq \ell \leq D$ , we now integrate over all lengths  $0 \leq \ell \leq D$  and all heights  $h_c, h_d$  with  $0 \leq h_c, h_d \leq \ell$  with respect to their corresponding densities. This yields

$$\begin{aligned} \mathbb{P}(\mathcal{E}_{p_a p_b, p_c, p_d}) &\leq \int_{\ell=0}^1 \int_{h_c=0}^{\ell} \int_{h_d=0}^{\ell} 2\pi\ell(3\ell)^2 \left(1 - \frac{\ell(h_c + h_d)}{2}\right)^{n-4} dh_d dh_c d\ell \\ &\quad + \int_{\ell=1}^D \int_{h_c=0}^{\frac{2}{\ell}} \int_{h_d=0}^{\frac{2}{\ell}} 2\pi\ell(3\ell)^2 \left(1 - \frac{\ell(h_c + h_d)}{2}\right)^{n-4} dh_d dh_c d\ell, \end{aligned}$$

and using  $1 - x \leq e^{-x}$ , we obtain

$$\begin{aligned} \mathbb{P}(\mathcal{E}_{p_a p_b, p_c, p_d}) &\leq 18\pi \int_{\ell=0}^1 \int_{h_c=0}^{\ell} \int_{h_d=0}^{\ell} \ell^3 \exp\left(-\frac{\ell(h_c + h_d)}{2}(n-4)\right) dh_d dh_c d\ell \\ &\quad + 18\pi \int_{\ell=1}^D \int_{h_c=0}^{\frac{2}{\ell}} \int_{h_d=0}^{\frac{2}{\ell}} \ell^3 \exp\left(-\frac{\ell(h_c + h_d)}{2}(n-4)\right) dh_d dh_c d\ell. \end{aligned}$$

Evaluating these two integrals separately, we get

$$\begin{aligned} \mathbb{P}(\mathcal{E}_{p_a p_b, p_c, p_d}) &\leq \frac{36\pi n}{(n-4)^3} + O\left(\frac{1}{n^3}\right) \\ &\quad + \frac{36(D^2-1)\pi}{(n-4)^2} + o\left(\frac{1}{n^2}\right) \\ &= \frac{36D^2\pi}{n^2} + o\left(\frac{1}{n^2}\right). \end{aligned}$$

Thus, since there are  $\binom{n}{2}$  choices for the points  $p_a$  and  $p_b$ , and at most  $n^2$  choices for the points  $p_c$  and  $p_d$ , we have

$$\mathbb{E}[C_4] \leq \binom{n}{2} n^2 \mathbb{P}(\mathcal{E}_{p_a p_b, p_c, p_d}) \leq 18D^2\pi n^2 + o(n^2),$$

proving the upper bound of the theorem. A lower quadratic bound is well known, see e.g. [7].  $\square$

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