

## RANDOM CENTRAL LIMIT THEOREMS FOR MARTINGALES

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**1. Introduction.** Suppose  $\{X_n\}$  is a strictly stationary ergodic process such that

$$(1) \quad \mathbf{E}(X_1) = 0, \mathbf{E}(X_n|X_1, \dots, X_{n-1}) = 0, n \geq 2$$

and

$$(2) \quad \mathbf{E}(X_1^2) = 1.$$

Let

$$(3) \quad \zeta_n = X_1 + X_2 + \dots + X_n, \quad \zeta_0 = 0;$$

and

$$(4) \quad \eta_n = n^{-\frac{1}{2}} \sum_{i=1}^n X_i.$$

It is known that  $\eta_n$  is asymptotically normal with mean 0 and variance 1 from a theorem of BILLINGSLEY [1]. Under a strong mixing condition on  $\{X_n\}$ , we shall now obtain a random version of this theorem similar to the theorems obtained by RÉNYI [5] and BLUM, HANSON, ROSENBLATT [2] for sums of independent random variables.

Suppose that  $\{v_n\}$  is a sequence of positive integer valued random variables such that  $\{v_n\}$  is independent of  $\{X_n\}$  and  $v_n \rightarrow +\infty$  in probability. Then, it follows, from the arguments in RÉNYI [5], that

$$(5) \quad \lim_{n \rightarrow \infty} \mathbf{P}(\eta_{v_n} \leq x) = \Phi(x),$$

where

$$(6) \quad \Phi(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt.$$

We shall now assume that the stationary process  $\{X_n\}$  satisfies the following modified version of the strong mixing condition of ROSENBLATT [7]. Let  $m_a^b$  denote the  $\sigma$ -field generated by the random variables  $X_n, a \leq n \leq b$ . Then

$$(7) \quad \sup_{A \in m_a^c, B \in m_{c+n}^\infty} |\mathbf{P}(B|A) - \mathbf{P}(B)| \leq \alpha(n) \quad \text{where } \alpha(n) \rightarrow 0 \text{ as } n \rightarrow \infty.$$

**THEOREM 1.** *Under condition (7), if  $\{v_n\}$  is a sequence of positive integer valued random variables such that  $n^{-1}v_n$  converges in probability to a positive random variable  $\mu$  having a discrete distribution, then (5) holds.*

**THEOREM 2.** Under condition (7), if  $\{v_n\}$  is a sequence of positive integer valued random variables such that  $n^{-1}v_n$   $\epsilon$ converges in probability to a positive random variable  $\mu$ , then (5) holds.

Note that theorem 1 is a special case of theorem 2. We shall prove theorem 1 and make use of it in the proof of theorem 2. Proofs of the above theorems follow the corresponding proofs of RÉNYI [5], BLUM, HANSON and ROSENBLATT [2] for sums of independent random variables.

**2. Some lemmas.** In this section, we shall state and prove some lemmas which will be used later.

**LEMMA 1.** If  $\{\tau_n\}$  is a sequence of random variables with  $\mathbf{E}(\tau_k|\tau_1, \dots, \tau_{k-1})=0$  and  $\sigma_k^2 = \text{var} [\tau_1 + \tau_2 + \dots + \tau_k]^2 < \infty$  for  $1 \leq k \leq n$ , then for any  $\epsilon > 0$

$$(8) \quad \mathbf{P}\left(\text{Max}_{1 \leq k \leq n} |\tau_1 + \tau_2 + \dots + \tau_k| > \epsilon\right) \leq \frac{\sigma_n^2}{\epsilon^2}.$$

Proof of this lemma can be found in DOOB [4].

**LEMMA 2.** Let  $W_n, X_{m,n}, Y_{m,n}^{(j)}$  and  $Z_{m,n}^{(j)}$  be random variables for  $m, n=1, 2, \dots$  and  $j=1, \dots, k$ . Suppose

$$W_n = X_{m,n} + \sum_{j=1}^k Y_{m,n}^{(j)} Z_{m,n}^{(j)}$$

and

- (a)  $\lim_{m \rightarrow \infty} \limsup_n \mathbf{P}(|Y_{m,n}^{(j)}| > \epsilon) = 0$  for every  $\epsilon > 0$  and  $1 \leq j \leq k$ ;
- (b)  $\lim_{M \rightarrow \infty} \limsup_m \limsup_n \mathbf{P}(|Z_{m,n}^{(j)}| > M) = 0$  for  $j=1, \dots, k$ ;
- (c) the distributions of  $\{X_{m,n}\}$  converge to the distribution function  $F$  for each fixed  $m$ .

Then the distribution function of  $\{W_n\}$  converges to  $F$ .

Proof of this lemma can be found in BLUM, HANSON and ROSENBLATT [2].

**DEFINITION.** A sequence of random variables  $\{\eta_n\}$  is said to be mixing in the sense of Rényi if for any event  $A$  with  $\mathbf{P}(A) > 0$ , and for any real number  $x$ ,

$$\lim_n \mathbf{P}(A_n|A) = \lim_n \mathbf{P}(A_n)$$

where  $A_n = [\eta_n \leq x]$ .

**LEMMA 3.** Under the strong mixing condition on the process  $\{X_n\}$ , the sequence  $\{\eta_n\}$  is mixing.

**PROOF.** Let  $A_k = (\eta_k \leq x)$ . By Theorem 2 of RÉNYI [6] it is enough to prove that for any  $A_k$ ,

$$\lim_n \mathbf{P}(A_n|A_k) = \lim_n \mathbf{P}(A_n).$$

Let  $j_n = [n^{1/4}]$  and consider the sequence of random variables  $n^{-1/2}\zeta_{k+j_n}$ . Since  $\mathbf{E}(\zeta_{k+j_n})=0$  and  $\text{Var}(\zeta_{k+j_n})=(k+j_n)$ ,  $n^{-1/2}\zeta_{k+j_n}$  converges to zero in probability as  $n$  tends to  $\infty$ . Therefore

$$(9) \quad \lim_n \mathbf{P}(A_n|A_k) = \lim_n \mathbf{P}(n^{-1/2}\{\zeta_n - \zeta_{k+j_n}\} \leq x|A_k).$$