

## MARTINGALE RANDOM CENTRAL LIMIT THEOREMS

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**1. Introduction.** Let  $\{S_n, \mathcal{F}, n \geq 1\}$  be a martingale on the probability space  $(\Omega, \mathcal{A}, P)$ , with  $S_0 = 0$ , and  $X_n = S_n - S_{n-1}$ ,  $n \geq 1$ .  $\mathcal{F}_0$  need not be the trivial  $\sigma$ -field  $\{\emptyset, \Omega\}$ . Let  $\varphi_j(t) = E(\exp \{itX_j\} | \mathcal{F}_{j-1})$ , and let  $\sigma_j^2 = E(X_j^2 | \mathcal{F}_{j-1})$ ,  $s_n^2 = \sum_{j=1}^n \sigma_j^2$  for  $n = 1, 2, \dots$ .

Now let  $\{N_n, n \geq 1\}$  be a sequence of positive integer-valued random variables defined on the same probability space  $(\Omega, \mathcal{A}, P)$ . Let us denote

$$S_{N_n} = X_1 + X_2 + \dots + X_{N_n}, \quad V_n^2 = \sum_{j=1}^{N_n} \sigma_j^2, \quad f_n(t) = \prod_{j=1}^{N_n} \varphi_j(t/B_n), \quad b_n = (\max_{k \leq N_n} \sigma_k^2) / B_n^2,$$

where  $B_n^2 = E(X_1^2 + X_2^2 + \dots + X_{N_n}^2)$ . Throughout the paper  $B_n$  is assumed to be finite for all  $n \geq 1$ .

In what follows  $C = C[0, 1]$  denotes the space of real-valued, continuous functions on  $[0, 1]$  and  $\mathcal{G}$  denotes a  $\sigma$ -field of Borel sets generated by the open sets of uniform topology. By  $W$  we will denote the Wiener measure on  $(C, \mathcal{G})$  with the corresponding Wiener process  $\{W(t): 0 \leq t \leq 1\}$ , (cf. [1], Sec. 9).

Let  $Y_n(t)$ ,  $0 \leq t \leq 1$ , be the random function defined as follows:

$$(1) \quad Y_n(t) = S_k/V_n + X_{k+1}(tV_n^2 - s_k^2)/V_n \sigma_{k+1}^2$$

for  $0 \leq t \leq 1$  and  $s_k^2 \leq tV_n^2 \leq s_{k+1}^2$ ,  $k = 0, 1, 2, \dots, N_n - 1$ , where  $s_0^2 = 0$ . It is obvious that  $Y_n(t)$  is continuous with probability one, being composed of straight line segments joining the points  $(s_k^2/V_n^2, S_k/V_n)$ ,  $k = 0, 1, 2, \dots, N_n$ . Thus there is a measure  $P_n$  in the space  $(C, \mathcal{G})$ , according to which the stochastic process  $\{Y_n(t), 0 \leq t \leq 1\}$  is distributed.

In this paper we use an approach developed by BROWN [2], to generate random central limit theorems for martingales. Section 2 defines the random Lindeberg condition for martingales and gives its several equivalent forms. Theorems 1 and 2 generalize Lindeberg—Feller's central limit theorem to random sums. Section 3 contains an invariance principle for a certain class of martingales. From this result we obtain some new limit theorems concerning of sums with random indices. The results obtained are generalizations of that of given in [5–7].

Throughout, we use the notations  $X_+ = \max(0, X)$  and  $X_- = \max(0, -X)$ , while  $\operatorname{Re} z$  is used to denote the real part of  $z$ .  $\Phi(x)$  denotes the standard normal distribution function, and the various kinds of convergence, in  $L_p$  norm, in probability, and weak (in distribution) are denoted by  $\xrightarrow{L_p}$ ,  $\xrightarrow{P}$  and  $\xrightarrow{\mathcal{D}}$ , respectively.

**2. Random central limit theorems.** Throughout the paper we say that the random Lindeberg condition is satisfied if, for all  $\varepsilon > 0$ ,

$$(2) \quad L(n, \varepsilon) = B_n^{-2} E \left\{ \sum_{i=1}^{N_n} X_i^2 I(|X_i| \geq \varepsilon B_n) \right\} \rightarrow 0 \quad \text{as } n \rightarrow \infty,$$

where  $I(A)$  denotes the indicator function of the set  $A$ . Furthermore, we consider martingales for which

$$(3) \quad V_n^2/B_n^2 \xrightarrow{P} 1 \quad \text{as } n \rightarrow \infty.$$

First we shall prove the following

LEMMA 1. *If  $N_n$  is stopping time with respect to  $\{\mathcal{F}_i, i \geq 0\}$ , then (3) is equivalent to*

$$(4) \quad V_n^2/B_n^2 \xrightarrow{L_1} 1 \quad \text{as } n \rightarrow \infty.$$

PROOF. It is obvious that (4) implies (3). On the other hand, for every  $\varepsilon > 0$ ,

$$E(V_n^2/B_n^2 - 1)_- \leq \varepsilon + P(|V_n^2/B_n^2 - 1| \geq \varepsilon).$$

Thus, by (3),  $\lim_{n \rightarrow \infty} E(V_n^2/B_n^2 - 1)_- = 0$ . Furthermore,

$$\begin{aligned} EV_n^2 &= \sum_{k=1}^{\infty} \int_{[N_n=k]} \sum_{j=1}^k \sigma_j^2 dP = \sum_{k=1}^{\infty} \int_{[N_n \geq k]} E(X_k^2 | \mathcal{F}_{k-1}) dP = \\ &= \sum_{k=1}^{\infty} \int_{[N_n \geq k]} X_k^2 dP = \sum_{k=1}^{\infty} \int_{[N_n=k]} \sum_{i=1}^k X_i^2 dP = E(X_1^2 + \dots + X_{N_n}^2) = B_n^2. \end{aligned}$$

Thus  $E(V_n^2/B_n^2 - 1) = 0$ , which implies (4) since  $E(V_n^2/B_n^2 - 1)_+ = E(V_n^2/B_n^2 - 1) \rightarrow 0$  as  $n \rightarrow \infty$ .

Let us put

$$g(n, \varepsilon) = V_n^{-2} \sum_{j=1}^{N_n} E(X_j^2 I(|X_j| \geq \varepsilon B_n) | \mathcal{F}_{j-1}), \quad G(n, \varepsilon) = V_n^2 B_n^{-2} g(n, \varepsilon),$$

$$h(n, \varepsilon) = V_n^{-2} \sum_{j=1}^{N_n} E\{X_j^2 U(|X_j| B_n^{-1} \varepsilon^{-1}) | \mathcal{F}_{j-1}\}, \quad H(n, \varepsilon) = V_n^2 B_n^{-2} h(n, \varepsilon),$$

where  $U(x)$  is any continuous nonnegative function of bounded variation on  $[0, \infty)$  for which  $U(0) = 0$  and  $U(x) \rightarrow \text{const. } (> 0)$  as  $x \rightarrow \infty$ .

LEMMA 2. *Assume that for every  $n$ ,  $N_n$  is a stopping time with respect to  $\{\mathcal{F}_i, i \geq 0\}$ . Then, under the condition (3), or alternatively (4), the random Lindeberg condition is equivalent to the convergence to zero as  $n \rightarrow \infty$  of  $g(n, \varepsilon)$ ,  $G(n, \varepsilon)$ ,  $h(n, \varepsilon)$  or  $H(n, \varepsilon)$ , for all  $\varepsilon > 0$ , either in probability or in  $L_1$ .*

PROOF. It suffices to show the mutual equivalence of convergences in probability of  $g$ ,  $G$ ,  $h$  and  $H$ , then to show that each such convergence in probability implies a corresponding convergence in  $L_1$  norm since the convergence in  $L_1$  is stronger than convergence in probability.